

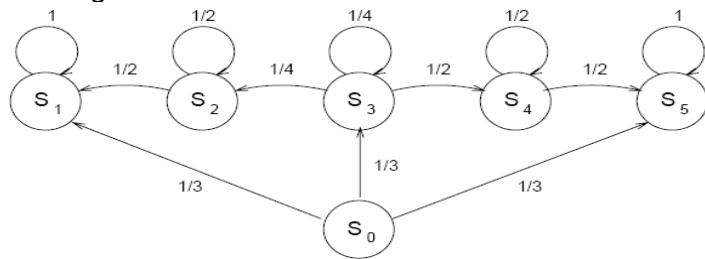
## Tutorial 6

- If “whether tomorrow is raining” is only related with the weather of today, and suppose the probability of “Both of today and tomorrow are raining” is 0.8; the probability of “Neither today nor tomorrow is raining” is 0.7. The state of raining or not on any given day is 1 and 0, respectively.

  - Is “whether it’s raining on any given day” a Markov chain. If yes, find the matrix of transition probability.
  - Find the probability of “Day 3” is raining, given “Day 1” is raining.
- Let  $X_1, X_2, \dots$  be independent, identical distributed random variables such that  $P[X_i = j] = \alpha_j$ ,  $j \geq 0$ . We say that a record occurs at time  $n$  if  $X_n > \max(X_1, \dots, X_{n-1})$ , where  $X_0 = -\infty$ , and if a record does occur at time  $n$ , let  $X_n$  be called the record value. Let  $R_i$  denote the  $i$ th record value.

  - Argue that  $\{R_i, i \geq 1\}$  is a Markov chain.
  - Compute its transition probabilities.

3. Consider the following Markov chain:



Given that the above process is in state  $S_0$  just before the first trial, determine by inspection the probability that:

- (1) The process enters  $S_2$  for the first time as the result of the  $k - th$  trial
- (2) The process enters  $S_2$  and then leaves  $S_2$  on the next trial.
- (3) The process enters  $S_1$  for the first time on the third trial.
- (4) The process is in state  $S_3$  immediately after the nth trial.

### Solution 6

#### 1. Solution:

(1) We define the state “Whether it’s raining on any given day” as  $X(n)$ , there are two states for  $X(n) : 0, 1$ .  $n$  is the day sequence.  $\{X(n), n \in \mathbb{Z}\}$  is a random sequence. From the definition of Markov process and the given question, this is a Markov chain. The corresponding matrix of transition probability is:

$$(P_{ij}(1)) = \begin{pmatrix} P_{00} & P_{01} \\ P_{10} & P_{11} \end{pmatrix} = \begin{pmatrix} 0.7 & 0.3 \\ 0.2 & 0.8 \end{pmatrix}$$

(2) By the Markov property, we have  $(P_{ij}(n)) = (P_{ij}(1))^n$ . Then,

$$(P_{ij}(3)) = \begin{pmatrix} 0.7 & 0.3 \\ 0.2 & 0.8 \end{pmatrix}^3 = \begin{pmatrix} 0.475 & 0.525 \\ 0.350 & 0.650 \end{pmatrix}$$

Because “Day 1” is raining, then  $\pi_1(n) = 1$  and  $\pi_0(n) = 0$ .

The state vector for “Day 3” is:

$$\begin{aligned} (\pi_0(n+3), \pi_1(n+3)) &= (\pi_0(n), \pi_1(n))(P_{ij}(3)) \\ &= (0,1) \begin{pmatrix} 0.475 & 0.525 \\ 0.350 & 0.650 \end{pmatrix} \\ &= (0.350, 0.650) \end{aligned}$$

Therefore, the probability of “Day 3” is raining, given “Day 1” is raining, is 0.650.

#### 2. Solution:

(1) Consider that the current record value is  $R_i$  for some  $i \geq 1$ . The value of the next record value  $R_{i+1}$  must be larger than  $R_i$ . So given  $R_i$ , the value of  $R_{i+1}$  is completely characterized by the  $X_i$ ’s, which are i.i.d. Therefore,  $\{R_i, i \geq 1\}$  is a Markov chain.

(2) Without loss of generality, let  $X_k$  be the element in the sequence  $\{X_j, j \geq 1\}$  that corresponds to  $R_{i+1}$ . We have:

$$P(R_{i+1} = n | R_i = m) = P(X_k = n | X_k > m) = \frac{\alpha_n}{\sum_{j=m+1}^{\infty} \alpha_j}$$

Where  $n > m$ . So, considering the cases with  $n \leq m$  leads to:

$$P(R_{i+1} = n | R_i = m) = \begin{cases} 0, & \text{for } n \leq m \\ \frac{\alpha_n}{\sum_{j=m+1}^{\infty} \alpha_j}, & \text{for } n > m \end{cases} \quad \text{for all } m=0,1,2,\dots$$

#### 3. Solution:

(1) Let  $A_k$  be the event that the process enters  $S_2$  for first time on trial  $k$ . The only way to enter state  $S_2$  for the first time on the  $k$ th trial is to enter state  $S_3$  on the first trial, remain in  $S_3$  for the next  $k-2$  trials, and finally enter  $S_2$  on the last trial. Thus,

$$P(A_k) = P_{03} \cdot P_{33}^{k-2} \cdot P_{32} = \frac{1}{3} \cdot \left(\frac{1}{4}\right)^{k-2} \cdot \frac{1}{4} = \frac{1}{3} \left(\frac{1}{4}\right)^{k-1}, \quad \text{for } k=2,3,\dots;$$

(2)

$$\begin{aligned} &P(\{\text{Process enters } S_2 \text{ and then leaves } S_2 \text{ on next trial}\}) \\ &= P(\{\text{Process enters } S_2\})P(\{\text{leaves } S_2 \text{ on next trial}\} | \{\text{in } S_2\}) \\ &= \left[ \sum_{k=2}^{\infty} P(A_k) \right] \cdot \frac{1}{2} = \left[ \sum_{k=2}^{\infty} \frac{1}{3} \left(\frac{1}{4}\right)^{k-1} \right] \cdot \frac{1}{2} = \frac{1}{6} \cdot \frac{1/4}{1-1/4} = \frac{1}{18} \end{aligned}$$

(3) This event can only happen if the sequence of state transitions is as follows:

$$S_0 \rightarrow S_3 \rightarrow S_2 \rightarrow S_1$$

Thus,  $P(\{\text{Process enters } S_1 \text{ for first time on third trial}\}) = P_{03}P_{32}P_{21} = \frac{1}{3} \cdot \frac{1}{4} \cdot \frac{1}{2} = \frac{1}{24}$ .

(4)

$$\begin{aligned} &P(\{\text{Process in } S_3 \text{ immediately after the } N \text{th trial}\}) \\ &= P(\{\text{moves to } S_3 \text{ in first trial and stays in } S_3 \text{ for next } N-1 \text{ trials}\}) \\ &= \frac{1}{3} \cdot \left(\frac{1}{4}\right)^{n-1}, \quad \text{for } n=1,2,3,\dots \end{aligned}$$